

# HMM for modeling aligned multiple sequences: phylo-HMM & multivariate HMM

Yuzhen Ye

School of Informatics and Computing

Indiana University, Bloomington

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# Content

- Consideration of two aligned sequences
    - TWINSKAN
  - Generalization to approaches for modeling multiple, aligned sequences
  - Phylo-HMM
    - Definition of phylo-HMM
    - Pruning algorithm for calculating the probability of a column (of the multiple alignment) given a phylogenetic model
  - Multivariate HMM
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# TWINSCAN model for gene finding in human and mouse genomes

- TWINSCAN is an augmented version of the GHMM used in Genscan.
- Input: syntenic regions in human and mouse genome
  - assumption: the gene structure (exon/intron boundaries) is *conserved* in these two genomes, and the conserved boundaries are aligned *precisely* in the pairwise genome alignment
- Output: the annotation of gene structure



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# TWINSCAN model

Human: **ACGGCGACUGUGCACGU**

Mouse: **ACUGUGAC GUGCACUU**

Alignment: **| | : | : | | | - | | | | | | : |**

intron exon

Why using multiple sequences?

- 1) multiple sequences gives strong signal (e.g. if a sequence profile of a splicing site is preserved, it is more likely to be a true splicing site); and more importantly,
  - 2) the conservation pattern can be used to discriminate exons and introns: exons tend to be more conserved than introns.
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# TWINSCAN algorithm

**The key idea:** converting a pairwise alignment into a single observation sequence on an expanded alphabet

1. Align the two sequences (e.g. from human and mouse genomes);
  2. Use the similar hidden states as Genscan;
  3. Design a **new “alphabet”** for observation symbols:  
4 x 3 = 12 symbols:  
 $\Sigma = \{ A-, A:, A|, C-, C:, C|, G-, G:, G|, U-, U:, U| \}$   
gap ( - ), mismatch ( : ), match ( | )
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# Example

Human: **ACGGCGACUGUGCACGU**

Mouse: **ACUGUGAC GUGCACUU**

Alignment: **||:|:| | - | | | | | : |**

Input to TWINSKAN HMM (observation sequence)

**A | C | G : G | C : G | A | C | U - G | U | G | C | A | C | G : U |**

Recall,  $e_E(\mathbf{A}|) > e_I(\mathbf{A}|)$  and  $e_E(\mathbf{A}-) < e_I(\mathbf{A}-)$

Likely exon will be annotated for the entire region

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# N-SCAN

- GHMM in TWINSCAN outputs a target genomic sequence and a **conservation sequence**
- **GHMM in N-SCAN outputs a target genomic sequence and N informant sequences**

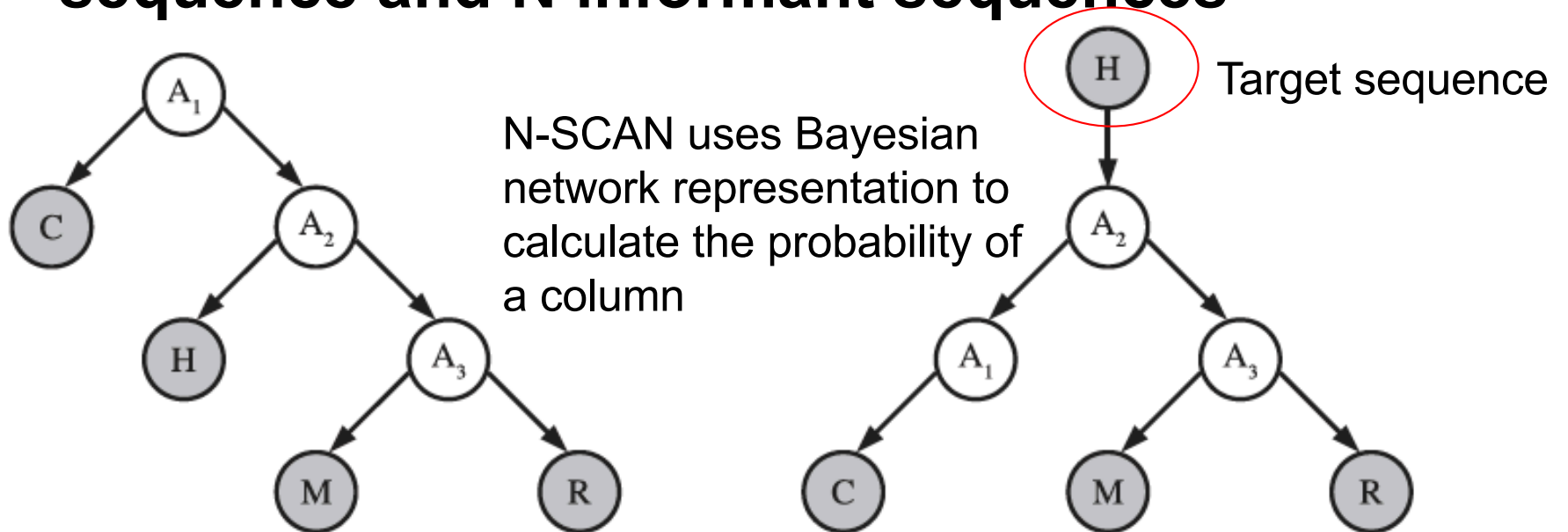


FIG. 1. A phylogenetic tree relating chicken (C), human (H), mouse (M), and rat (R). The graph can also be interpreted as a Bayesian network (left). The result of transforming the Bayesian network (right).

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# HMM for multiple aligned sequences

- Strategy 1: converting the alignment of multiple observation sequences into *one* observation sequence
    - New alphabet representing *convoluted* observation symbols, e.g., the TWINSCAN model
    - Not practical for  $n$  sequences: the size of the alphabet grows exponentially with  $O(2^n)$ .
  - Strategy 2: employing another probabilistic model to emit multiple aligned observation sequences simultaneously (Phylo-HMM model)
  - Strategy 3: emitting multiple aligned observation sequences simultaneously but independently, each following a different emission probability distribution (multivariate HMM)
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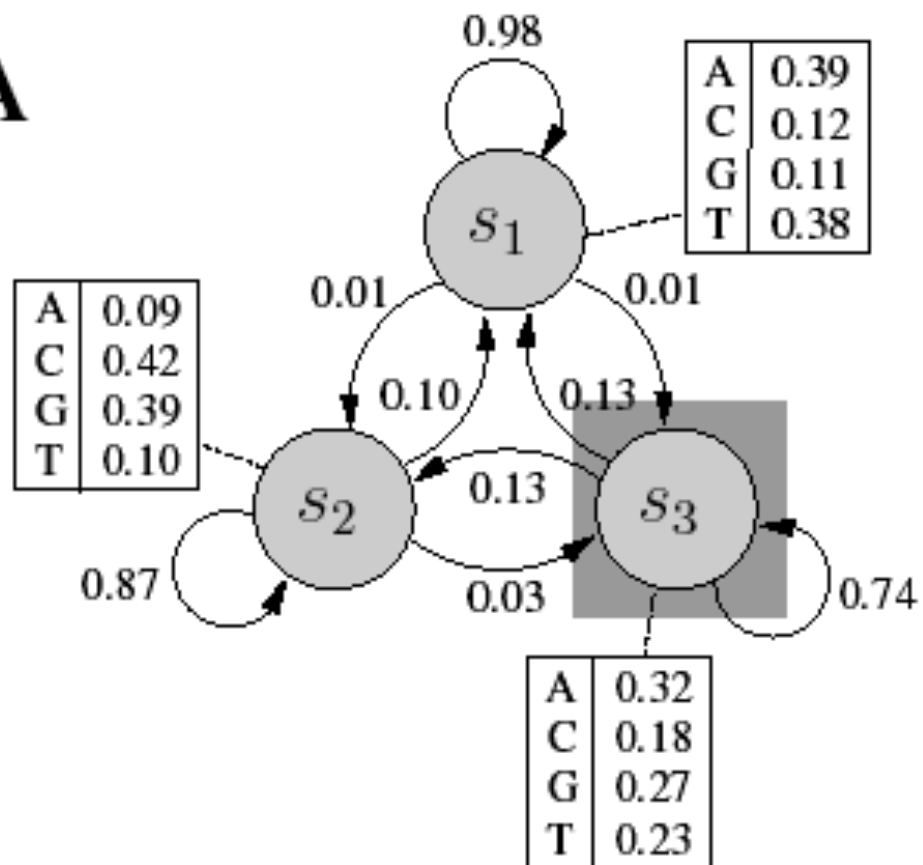
# Phylo-HMMs: model multiple alignments of syntenic sequences

- A phylo-HMM is a probabilistic machine that generates a multiple alignment, column by column, such that each column is defined by a phylogenetic model
  - Unlike single-sequence HMMs, the “emission” probabilities of phylo-HMMs are complex distributions defined by **phylogenetic models**
  - Molecular evolution can be viewed as a combination of two Markov processes
    - One operates in the dimension of **space** (along a genome)
    - One operates in the dimension of **time** (along the branches of a phylogenetic tree)
  - Phylo-HMMs combine phylogeny and HMM
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# Single-sequence HMM

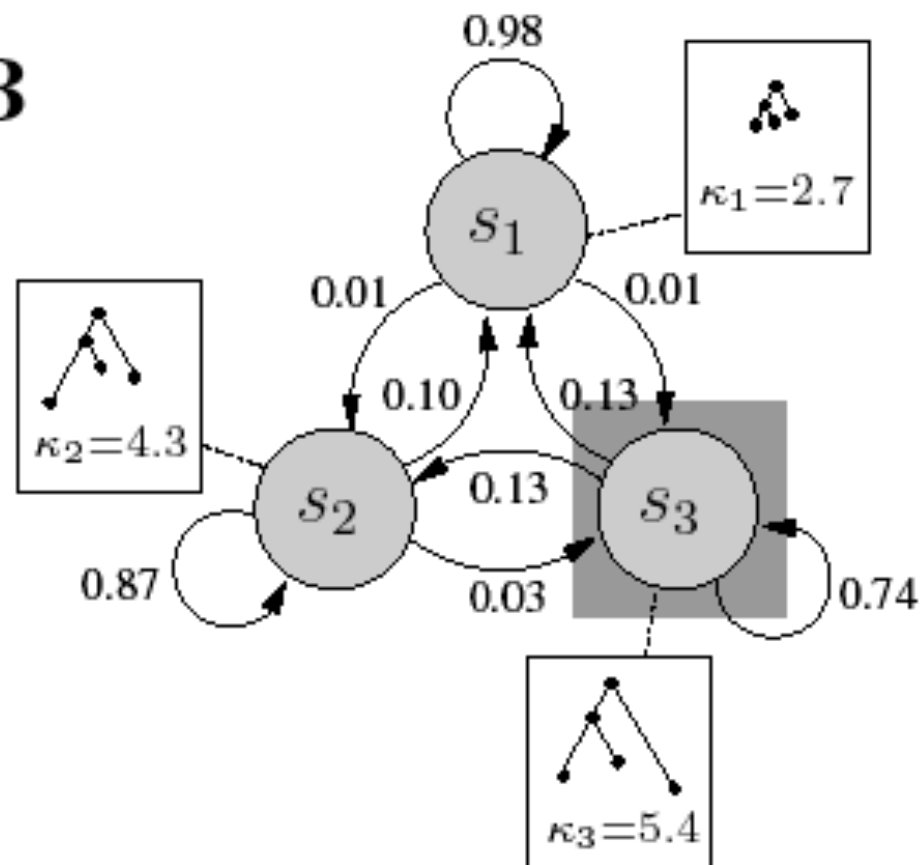
# Phylo-HMM

**A**



$X = \text{TAACGGCAGA} \dots$

**B**



$X = \begin{matrix} \text{TAACGGCAGA} \dots \\ \text{TTAGGCAAGG} \dots \\ \text{AAGGCGCCGA} \dots \end{matrix}$

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# Phylo-HMMs: formal definition

A phylo-HMM can be specified as  $\theta = (S, \psi, A, b)$ ,

- 1)  $S = \{S_1, S_2, \dots, S_M\}$ , a set of states
- 2)  $\psi = \{\psi_1, \psi_2, \dots, \psi_M\}$ , a set of associated phylogenetic models
- 3)  $A = \{a_{jk}\} (1 \leq j, k \leq M)$ , a matrix of state-transition probabilities
- 4)  $b = (b_1, \dots, b_M)$ , a vector of state-initial probabilities

$a_{jk}$  is the conditional probability of visiting state  $k$  at some site  $i$  given that state  $l$  is visited at site  $i - 1$ .  $b_j$  is the probability that state  $j$  is visited first.

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# Questions we can ask using phylo-HMM

A path through the phylo-HMM is a sequence of states  $\phi = (\phi_1, \dots, \phi_M)$ , such that  $\phi_i \in \{1, \dots, M\}$  for all  $1 \leq i \leq L$ .

The joint probability of a path and an alignment is,

$$p(\phi, X|\theta) = b_{\phi_1} P(X_1|\psi_{\phi_1}) \prod_{i=2}^L a_{\phi_{i-1}, \phi_i} P(X_i|\psi_{\phi_i})$$

The probability of the observation (likelihood) is,

$$p(X|\theta) = \sum_{\phi} P(\phi, X|\theta)$$

The probability of emitting column  $i$  given a **phylogenetic model**

Forward algorithm

The most probable (maximum-likelihood) path,

$$\hat{\phi} = \arg \max_{\phi} P(\phi, X|\theta)$$

Viterbi algorithm

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# Phylogenetic models

- The different phylogenetic models associated with the states of a phylo-HMM may **reflect different overall rates of substitution** (e.g. in conserved and non-conserved regions), different patterns of substitution or background distributions, or even different tree topologies (as with recombination)



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# Phylogenetic models

$$\psi_j = (Q_j, \pi_j, \tau_j, \beta_j)$$

HKY model

$Q_j$  : substitution rate matrix

$\pi_j$  : background frequencies

$\tau_j$  : binary tree

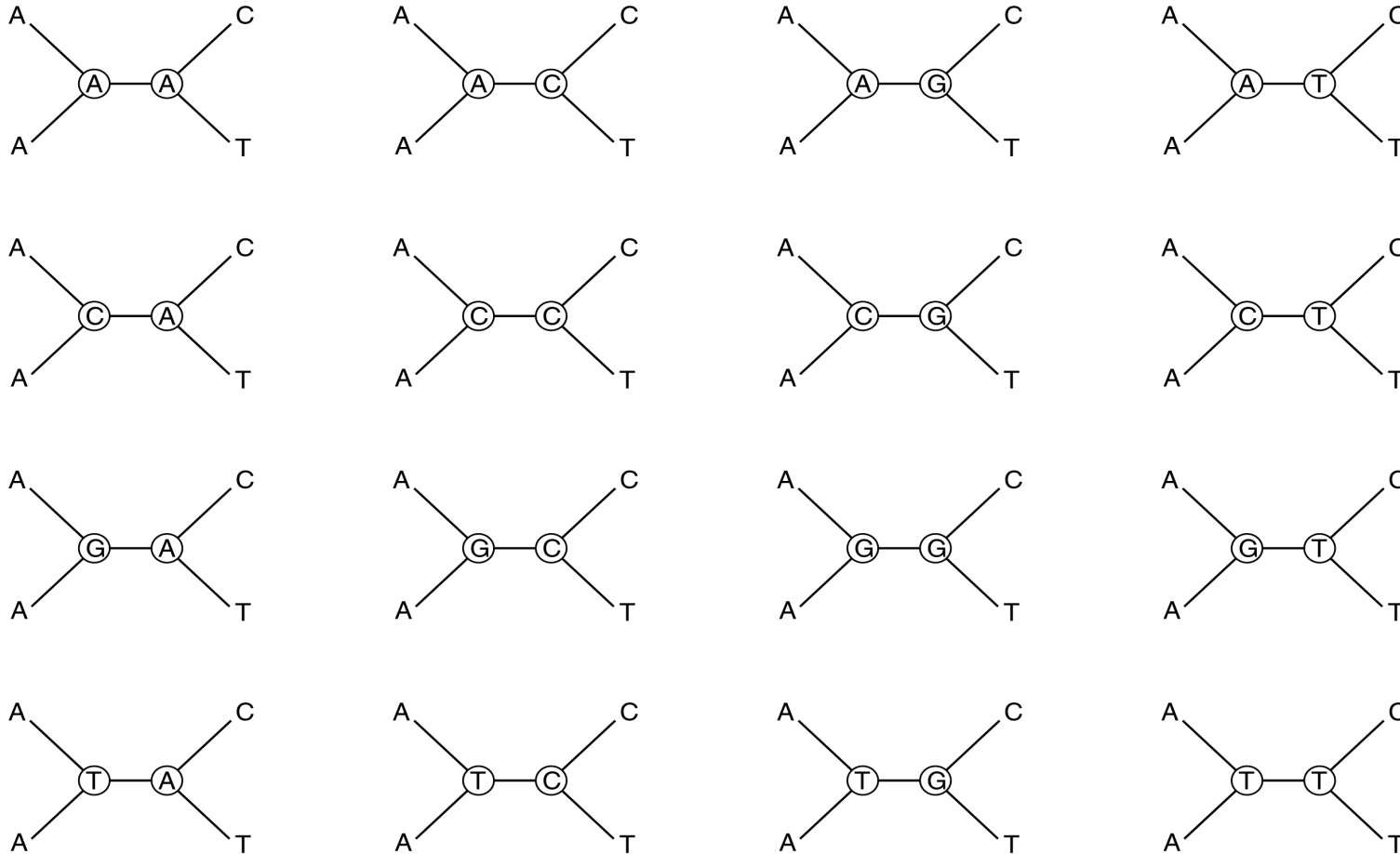
$\beta_j$  : branch lengths

$$Q_j = \begin{pmatrix} - & \pi_{C,j} & \kappa_j \pi_{G,j} & \pi_{T,j} \\ \pi_{A,j} & - & \pi_{G,j} & \kappa_j \pi_{T,j} \\ \kappa_j \pi_{A,j} & \pi_{C,j} & - & \pi_{T,j} \\ \pi_{A,j} & \kappa_j \pi_{C,j} & \pi_{G,j} & - \end{pmatrix}$$

- The model is defined with respect to an alphabet  $\Sigma$  of size  $d$
  - The substitution rate matrix has dimension  $d \times d$
  - The background frequencies vector has dimension  $d$
  - The tree has  $n$  leaves, corresponding to  $n$  extant taxa in the multiple alignment of observation sequences
  - The branch lengths are associated with the tree
  - *How to calculate the likelihood of a column given a model?*
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# Brute force approach to likelihood calculation

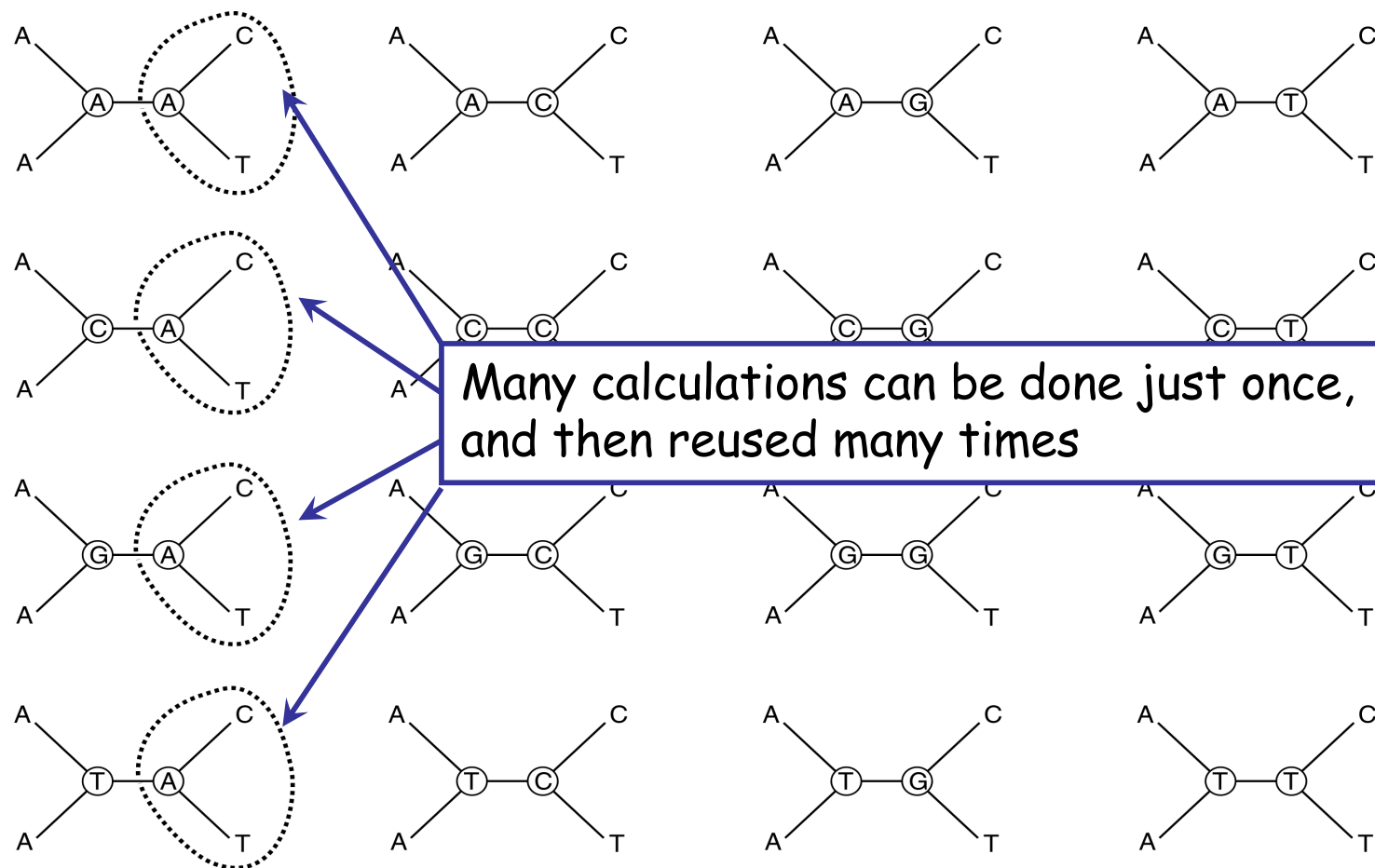


Given a column AACT

Brute force strategy: Try all 16 combinations of ancestral states and sum

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# Pruning algorithm



\*The pruning algorithm was introduced by: Felsenstein, J. 1981. Evolutionary trees from DNA sequences: a maximum likelihood approach. *Journal of Molecular Evolution* 17:368-376



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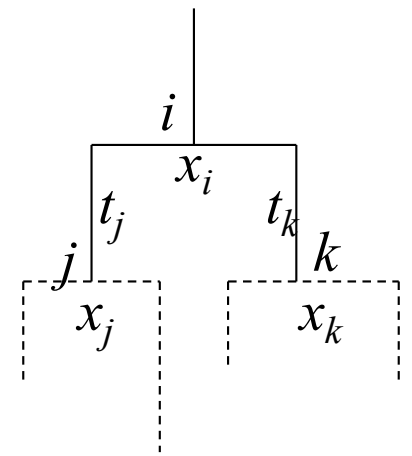
# Pruning algorithm

$P(X|\Psi)$  ( $X$ : a column;  $\Psi$ : phylogenetic model; subscript not shown for clarity) can be computed recursively from leaves to root.

$L_i(x_i)$ : the probability of observing leaves in the subtree *rooted* by  $i$ , while the root is assigned to  $x_i$  (A, T, C or G),

$$L_i(x_i) = \left\{ \sum_{x_j} p_{x_i x_j}(t_j) L_j(x_j) \right\} \times \left\{ \sum_{x_k} p_{x_i x_k}(t_k) L_k(x_k) \right\}$$

Sum over all possible  $x_j$  (A, T, C, or G)



$j$  &  $k$ : offspring nodes of node  $i$

$t_j$  &  $t_k$ : the branch lengths

$p_{a,b}(t)$ : the probability to observe a substitution from  $a$  to  $b$  within the evolution time of  $t$ .

The total probability at the root node  $r$ :  $P(X|\psi) = \sum_{x_r} L_r(x_r)$

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# Substitution probabilities

- Pruning algorithm requires the conditional probabilities of substitution  $p_{a,b}(t)$  for all bases  $a, b \in \Sigma$  and branch lengths  $t \in \beta$

Another way to denote this probability:  $P(b | a, t, \psi)$

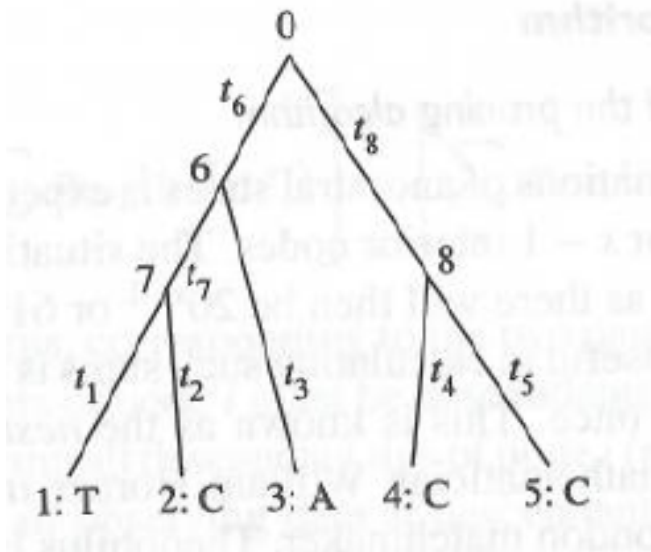
- It can be computed using a **continuous-time Markov model** of substitution, defined by the rate matrix  $Q$

$$P(t) = (e^Q)^t = e^{Qt}$$

(matrix multiplication)

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# Example



Substitution matrices:  $P(t=0.1)$   $P(t=0.2)$

$$P(0.1) = \begin{bmatrix} \text{T} & \text{C} & \text{A} & \text{G} \\ 0.906563 & 0.045855 & 0.023791 & 0.023791 \\ 0.045855 & 0.906563 & 0.023791 & 0.023791 \\ 0.023791 & 0.023791 & 0.906563 & 0.045855 \\ 0.023791 & 0.023791 & 0.045855 & 0.906563 \end{bmatrix}$$

$$P(0.2) = \begin{bmatrix} 0.825092 & 0.084274 & 0.045317 & 0.045317 \\ 0.084274 & 0.825092 & 0.045317 & 0.045317 \\ 0.045317 & 0.045317 & 0.825092 & 0.084274 \\ 0.045317 & 0.045317 & 0.084274 & 0.825092 \end{bmatrix}$$

Branch lengths:  $t_1=t_2=t_3=t_4=t_5=0.2$ ;  $t_6=t_7=t_8=0.1$ .

$$P(X_i | \psi_j) = P(TCACC | T, t_1, t_2, t_3, t_4, t_5, t_6, t_7, t_8, \kappa)$$

$$= \sum_{x_0} \sum_{x_6} \sum_{x_7} \sum_{x_8} \left[ \pi_{x_0} P_{x_0 x_6}(t_6) P_{x_0 x_8}(t_8) P_{x_6 x_7}(t_7) P_{x_7 T}(t_1) P_{x_7 C}(t_2) P_{x_6 A}(t_3) P_{x_8 C}(t_4) P_{x_8 C}(t_5) \right]$$

# Example

By using pruning algorithm, it can be computed through  $L_i(x_i)$ , from leaves to root.

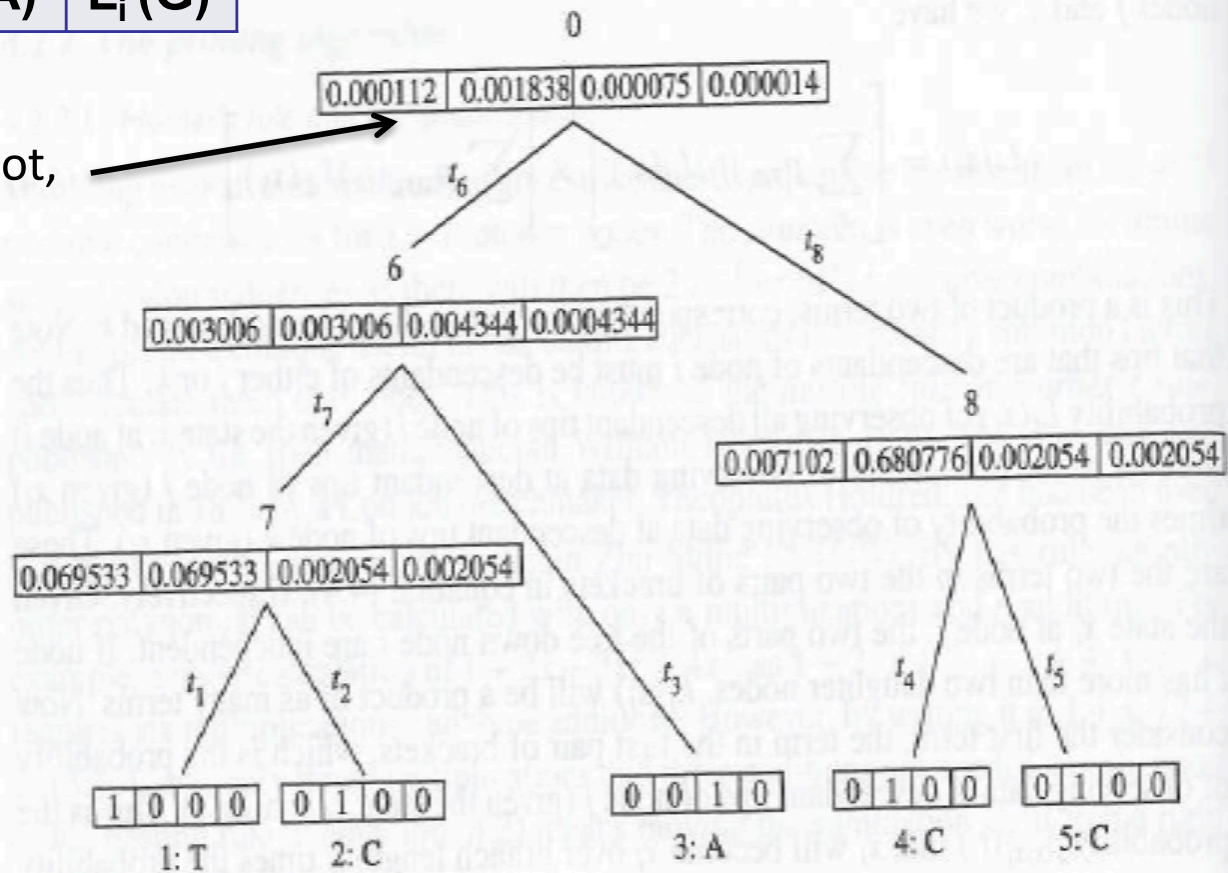
$L_i(T)$	$L_i(C)$	$L_i(A)$	$L_i(G)$
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P(TCACC)?

Summing at the root,

$P=0.000509843$

Initialization



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# Applications of Phylo-HMMs

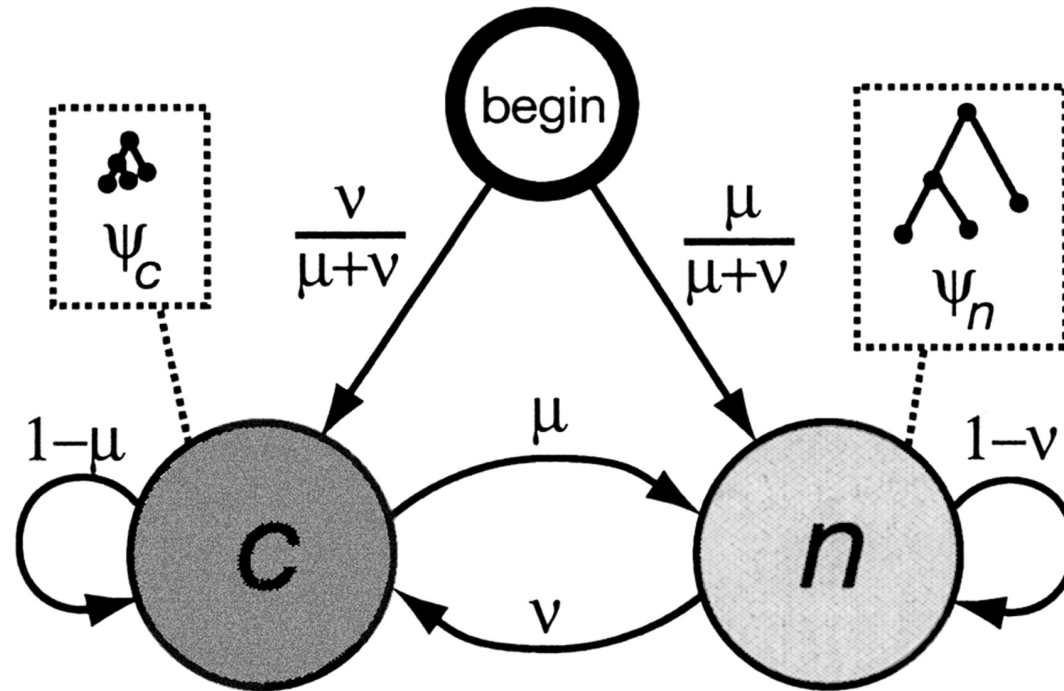
- Improving phylogenetic modeling that allow for variation among sites in the rate of substitution (Felsenstein & Churchill, 1996; Yang, 1995)
  - Protein secondary structure prediction (Goldman et al., 1996; Thorne et al., 1996)
  - Detection of recombination from DNA multiple alignments (Husmeier & Wright, 2001)
  - **Comparative genomics (Siepel, et. al. Haussler, 2005)--phastCons**
  - Inferring sequence regions under functional divergence in duplicate genes (Huang & Golding, Bioinformatics, 2012)
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# phastCons

- phastCons is based on a two-state phylogenetic hidden Markov model (phylo-HMM), with a state for conserved regions and a state for nonconserved regions; the free parameters of the model were estimated from a multiple alignment by maximum likelihood, using an EM algorithm.
  - Developed for searching for conserved elements in vertebrate genomes, using genome-wide multiple alignments of five vertebrate species (human, mouse, rat, chicken, and *Fugu rubripes*)
    - The predicted (conserved) elements cover roughly 3%–8% of the human genome (depending on the details of the calibration procedure)
    - HCEs (highly conserved elements) are associated with the 3' UTRs of regulatory genes, stable gene deserts, and megabase-sized regions rich in moderately conserved noncoding sequences. Noncoding HCEs also show strong statistical evidence of an enrichment for RNA secondary structure.
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State-transition diagram for the phylo-HMM used by phastCons, which consists of a state for conserved regions (c) and a state for nonconserved regions (n).



**x** = TCGCGACATATACGA . . .   
TTGGGGGCATGTGGGT . . .   
AGCAGACGTCCGCAA . . .

Siepel A et al. *Genome Res.* 2005;15:1034-1050



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# PHAST & RPHAST

- **PHAST and RPHAST: phylogenetic analysis with space/time models** ([Brief Bioinform.](#) 2011)
  - <http://compgen.bscb.cornell.edu/phast/>
  - <http://compgen.bscb.cornell.edu/rphast/>
- Include phastCons, phastOdds, phyloP, dbless, etc

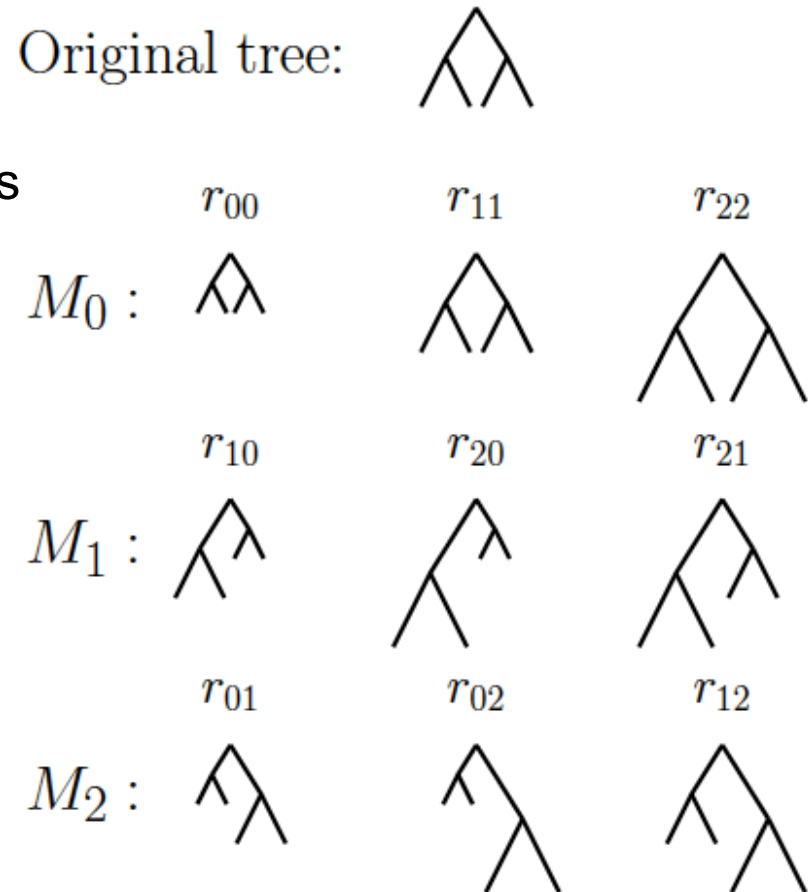




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# HMMDiverge

Inferring Sequence Regions under  
Functional Divergence in Duplicate Genes  
*Bioinformatics (2011)*

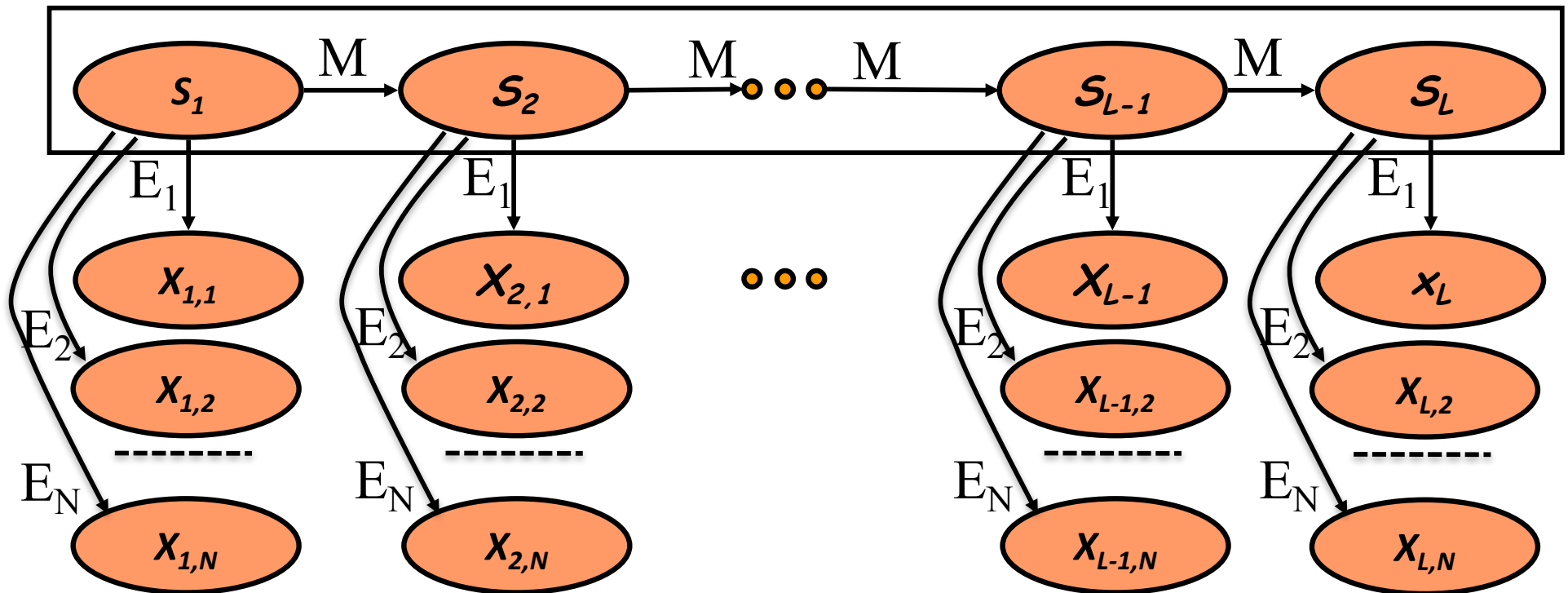


An example with three states:  
M0 (no functional divergence)  
M1 & M2 (with functional divergence)

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# Multivariate HMM



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# Multivariate HMM (formal definition)

- A multivariate HMM  $\theta$  has
    - **N** sets of observation symbols, each for one given observation sequence  $n$  ( $n=1, 2, \dots, N$ )
    - A set of hidden states
    - Transition probabilities  $a_{ij}$ , for any pair of hidden states  $i$  and  $j$
    - Initial probabilities  $B_j = a_{0j}$  for any hidden states  $j$
    - **N** sets of emission probabilities  $e_s^n(x_n)$  for the observation symbol being emitted in the  $n$ th observation sequence from the hidden state  $s$ .
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## Multivariate HMM

- Given  $N$  observation sequences of the same length  $L$ ,  $X = \{(x_{1,1} \dots x_{1,L}), \dots, (x_{N,1} \dots x_{N,L})\}$  and the hidden state sequence  $S = (s_1 \dots s_L)$ , the full probability from the multivariate HMM is,

$$P(S, X | \theta) = \prod_{j=1}^L \left[ a_{s_{j-1}s_j} \prod_{n=1}^N e_{s_j}(x_{n,j}) \right]$$

Let  $e_{s_i}(x_{n,1}, \dots, x_{n,j}) = \prod_{k=1}^j e_{s_i}(x_{n,k})$ , the multivariate HMM can be reduced to conventional HMM, except the observation symbol becomes a vector  $(x_{n,1} \dots x_{n,j})$  at position  $j$ . The same algorithms for model inference (Viterbi and forward/backward) and learning can be used.

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